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ABSTRACT

The choice of space and time windows is often a puzzling task in the simultaneous comparison of satellite and in situ observations. In this study, numerical simulations are carried out to examine the impacts of instrument noises, geophysical variabilities and the number of data pairs on the determination of an optimal size of space and time windows in the calibration/validation of altimeter-derived wind speeds against buoy measurements. Useful results are obtained which not only provide guidelines for future altimeter/buoy wind speed comparisons, but also help to explain and clarify some of the historical confusions and contradictions on the subject concerned.

1 Introduction

It is a common practice in remote sensing that satellite observations be calibrated and/or validated by field measurements. In compiling a colocated satellite/in situ dataset, the choice of space and time windows is very often a puzzling task because of the conflicting requirements involved. On the one hand, large windows are desirable for the inclusion of more colocation data in order to obtain low-variance-correlation statistics. On the other hand, however, small windows are expected for the exclusion of less coincident measurements in order to avoid additional variability due to the inhomogeneous distribution of the geophysical parameters. This problem could be more serious when dealing with sensors whose measurements are made on point basis, such as satellite altimeters. Since the overflight of an altimetric satellite rarely coincides with the field measurement either in location or in time, there almost always exist some spatial and temporal lags. It is therefore necessary to define some criteria for paring the measurements from two independent sources.

As far as the verification of altimeter-derived sea surface wind speed is concerned, the above-mentioned problem has been aware by the altimetric community for many years. Table 1 is a brief summary of the space and time windows used by previous investigators for comparison purposes. In the space domain, the size of the window ranges from 10 km to 150 km; While in the time domain, it varies from 0.5 h to 1.5 h. Obviously, there is a wide consideration among various authors on the definition of data colocation. Attempts have been made by a number of researchers to investigate the influences of window size on the comparison results. Their conclusions, however, are somewhat divided. One argument suggests that the variations of the space and time windows do not have a crucial impact on the accuracy of the satellite/buoy comparisons (Dobson et al., 1987; Glazman and Pilorz, 1990). In contrast, another argument suggests that reducing the spatial lag persistently improves the agreement of wind speed measurements from altimeters and ocean buoys, while time lags of up to 1 hour do not produce significant changes in the key statistics, such as the RMS difference and the correlation coefficients (Gower, 1996; Hwang et al., 1998). Some of those who favour the second argument go on to further point out that, for future calibration purposes, it is more critical to select in situ measurements that are closer to the satellite tracks while the temporal lag can be relaxed (Hwang et al., 1998).

Table 1. A partial list of the space and time windows used by previous investigators

for altimeter/buoy wind speed comparisons.

Reference	Satellite	Buoy	Space Window	Time Window
Brown et al.	GEOS-3	NOAA/NDBC	110 km	1.5 h

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(1981)		(USA)		
Dobson et al.	Geosat	NOAA/NDBC	50, 100, 150 km	0.5 h
(1987)		(USA)		
Monaldo & Dobson	Geosat	NOAA/NDBC	50 km	0.5 h
(1989)		(USA)		
Glazman & Pilorz	Geosat	NOAA/NDBC	0.25°, 1.0°	0.5, 1.0 h
(1990)		(USA)		
Witter & Chelton	Geosat	NOAA/NDBC	50 km	0.5 h
(1991)		(USA)		
Ebuchi et al.	Geosat	JMA	50, 100, 150 km	0.5, 1.0, 1.5 h
(1992)		(Japan)		
Carter et al.	Geosat	NOAA/NDBC	0.25°	1.0 h
(1992)		(USA)		
Glazman & Greysukh	Geosat	NOAA/NDBC	0.5°	0.75 h
(1993)		(USA)		
Gower	TOPEX/	AECEC &	10, 20, 107 km	0.5 h
(1996)	Poseidon	DFO (Canada)		
Hwang et al.	TOPEX/	NOAA/NDBC	10, 50 km	0.5, 1.0 h
(1998)	Poseidon	(USA)		

Before making any comment on the above arguments, it is helpful to recall the major error sources that cause altimeter and buoy estimates of wind speed to differ. These are (1) buoy instrument inaccuracies, (2) altimeter-related uncertainties, (3) spatial separation, and (4) temporal separation. Keeping in mind these factors, it is somehow less surprise to see the inconsistent conclusions drawn by various authors. As listed in Table 1, the altimeters and buoys employed are not the same, the times and locations as well as the marine conditions under which the measurements were taken could also be dramatically different. Diverse conclusions should therefore be expected.

In the present work, numerical experiments are carried out to study the influences of the size of space and time windows on the altimeter/buoy wind speed comparisons. As can be seen later on, our simulation results will, to a large extent, help to explain the observed inconsistencies over the problem concerned.

2 Scheme and Experiments

In this section, the Monte Carlo technique used in the numerical experiments of this study will be described. The considerations on the altimeter and buoy instrument noises as well as the wind speed variability model will be discussed.

2.1 Generation of a Gaussian distribution using Monte Carlo method

It is understood that the error distributions of geophysical measurements usually take the Gaussian form. This is presumably the case for wind speed (Monaldo, 1988). Assuming that $\tilde{x}_1, \tilde{x}_2, \dots, \tilde{x}_n$ are independent random variables with a uniform probability density function (pdf) between [-1,1]. According to the central limit theorem, a Gaussian distribution, \tilde{y} , with a variance of σ^2 and a mathematical expectation of μ can be approximated by

$$\widetilde{y}_{j} = \sqrt{\frac{3}{n}} \sigma \sum_{i=1}^{n} \widetilde{x}_{ij} + \mu$$
⁽¹⁾

when *j* gets large and *n* is appropriately chosen. For the case of oceanic wind speed, let $\sigma = 1.0$ m/s and $\mu=10.0$ m/s, the result of Monte Carlo simulation for n=10 and $j_{max}=10,000$ is shown in Figure 1. The theoretical probability density function of a Gaussian distribution

$$y(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left[-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right]$$
(2)

is also superimposed for comparison. It is apparent that the simulated distribution agrees very well with the theoretical prediction, indicating that the pseudo Gaussian distribution generated by Monte Carlo method is a good approximation to the theory.

2.2 Simulation scheme

As stated earlier, the factors that complicating the altimeter/buoy wind speed comparisons include the instrument noises of both sensors and the spatial and temporal lags in obtaining the measurements. In our simulations, we will first consider the impact of the space lag together with the sensor noises. The time lag can be treated in parallel. For simplicity while without losing generality, we consider a geometry such that the location of a buoy station coincides with that of an altimeter nadir point which lies on the only ground track entering the colocation window. Assuming a random variable, \tilde{U}_T , which represents the true

wind speed value at the buoy site. Let \widetilde{U}_A and \widetilde{U}_B be the corresponding wind speeds measured by the altimeter and the buoy respectively, we can then write

$$\left(\widetilde{U}_{A}\right)_{n} = \left(\widetilde{U}_{T}\right)_{n} + \left(\widetilde{\varepsilon}_{A}\right)_{n} \tag{3a}$$

$$\left(\widetilde{U}_B\right)_n = \left(\widetilde{U}_T\right)_n + \left(\widetilde{\varepsilon}_B\right)_n \tag{3b}$$

for the *n*th comparison, where $\tilde{\varepsilon}_A$ and $\tilde{\varepsilon}_B$ are the Gaussiandistributed instrument noises of the altimeter and the buoy, respectively.



Figure 1. The Gaussian probability density function of wind speed with a variance of σ =1.0 m/s and a mean value of μ =10.0 m/s generated by Monte Carlo simulation (the noisy curve). The number of random samples is 10,000. The theoretical Gaussian probability density function is also superimposed (the smooth curve).

At this point, we introduce the space variability model of the wind

speed, V. The altimeter-measured wind speed of the *m*th neighboring sample with respect to the buoy site can be expressed by

$$(\widetilde{U}_A)_{mn} = (\widetilde{U}_T)_n + (V)_m + (\widetilde{\varepsilon}_A)_n$$
(4)
Consider a space window which contains Multimeter measurements, the averaged wind speed is then given by

Consider a space window which contains M altimeter measurements, the averaged wind speed is then given by

$$(\widetilde{U}_{A}(M))_{n} = \frac{1}{M} \sum_{m=1}^{M} \left[(\widetilde{U}_{T})_{n} + (V)_{m} + (\widetilde{\varepsilon}_{A})_{n} \right]$$
(5)

We further assume that the wind speed variability model, V, has several possible modes (denoted by index I) corresponding to different oceanic conditions, and the altimeter noise, $\tilde{\mathcal{E}}_A$, has several possible levels (denoted by index J) corresponding to various generations of altimeters, Equation (5) becomes

$$(\widetilde{U}_{A}(I,J,M))_{n} = \frac{1}{M} \sum_{m=1}^{M} \left[(\widetilde{U}_{T})_{n} + (V(I))_{m} + (\widetilde{\varepsilon}_{A}(J))_{n} \right]$$
(6)

Ultimately, we examine the influence of the window size, M, under given conditions of V(I) and $\widetilde{\mathcal{E}}_A(J)$ through the following quantity,

$$D^{2}(I,J,M,N) = \frac{1}{N} \sum_{n=1}^{N} \left[(\widetilde{U}_{A}(I,J,M))_{n} - (\widetilde{U}_{B})_{n} \right]^{2}$$
(7)

where N is the total times of simulations. Inserting Equations (3b) and (6) into (7), yields,

$$D^{2}(I,J,M,N) = \frac{1}{N} \sum_{n=1}^{N} \left\{ \frac{1}{M} \sum_{m=1}^{M} \left[(\widetilde{U}_{T})_{n} + (V(I))_{m} + (\widetilde{\varepsilon}_{A}(J))_{n} \right] - \left[(\widetilde{U}_{T})_{n} + (\widetilde{\varepsilon}_{B})_{n} \right] \right\}^{2}$$
(8)

2.3 Numerical experiments

Before starting the numerical experiments, the values of the four parameters in Equation (8), I, J, M and N, must be properly selected, since they will, to a large extent, determine the validity and representativeness of the simulation results.

According to Monaldo (1988), the expected RMS differences between two wind speed estimates separated by a given distance or a given time appear to have a logarithmic form. The following wind speed variability models are employed in the space and time domain, respectively,

$$V^{s}(I) = I \times \ln(x/150 + 1)$$
 (9a)
 $V^{t}(I) = I \times \ln(t/1.5 + 1)$ (9b)

where x and t are spatial and temporal separations in kilometer and hour, respectively. The graphic illustrations of Equations (9a) and (9b) with I=0,1,2,3,4 (m/s) are presented in Figure 2. It is obvious that a greater I is associated with a higher wind speed variation.

As far as the parameter J is concerned, we set values of 1.2m/s, 0.8m/s, 0.4m/s and 0.0m/s to it, representing the decreasing noise levels of the past, present and future altimeters.

For parameter M, we consider a changeable window size varying from 0 to 150km in space and from 0 to 1.5h in time, which, to our knowledge, covers the maximum space and time windows used by previous investigators (see Table 1).



Figure 2. The assumed wind speed variabilities as functions of spatial and temporal lags for various *Is* (in m/s).

For parameter N, we will first make some initial tests on its effectiveness (see the next section), then, taking into account both the statistical confidence and the computational cost, a fixed value of N=10,000 will be used.

With the simulation scheme described early on, the numerical experiments can now proceed with a variety of combinations of I, J, M and N.

3 Results and Discussions

3.1 Influence of the number of colocation data

The first experiment is designed to examine the relationship between the expected RMS difference of the altimeter/buoy wind speed estimates and the number of colocation data used in the comparison. Consider a simple case of I=J=M=0, meaning that the wind field is uniform, the altimeter is perfect and only the sample at the buoy site is included. Consequently, the only source of error comes from the buoy. Select $\sigma_B=1.0$ m/s and $\mu_B=10.0$ m/s for $\tilde{\epsilon}_B$, the simulated D as a function of N is shown in Figure 3. As can be clearly seen from this figure, the RMS difference of wind speed decreases monotonically as the number of data pair increases. It is known from the theory of statistics that

$$(\boldsymbol{\sigma}_B)_N = \frac{\boldsymbol{\sigma}_B}{\sqrt{N}} \tag{10}$$

By overlaying the numerically simulated and theoretically predicted results at the bottom of Figure 3, the excellent agreement becomes immediately visible. This agreement serves as a direct verification on the validity of our numerical scheme.

3.2 Influence of the altimeter noise level

In order to test the impact of the altimeter errors on the choice of optimal window size, we assume that there are four generations of altimeters whose instrument noises are

Figure 3. The simulated RMS deviations of altimeter-measured wind speed as a function of the number of collocation data, N (the noisy curve at the bottom). A Gaussian pdf with σ =1.0 m/s and μ =10.0 m/s is employed in the simulation. A theoretical Gaussian pdf with the same properties is superimposed for comparison (the smooth curve at the bottom). The original biases of the simulated wind speed around the mean are also plotted as a function of N with an offset of 1 m/s.

at J=1.2, 0.8, 0.4, 0.0(m/s). Their corresponding relationships of simulated wind speed RMS with respect to spatial and temporal lags are plotted in Figures 4(a)-(d), respectively. Within each family, five modes of the wind speed variability model, corresponding to I=0, 1, 2, 3, 4(m/s), are applied.

The series of plots demonstrate a general trend that the optimal size of the space and time windows decreases monotonically with the noise level of the altimeters. Figures 4(a) and 4(d) represent two extreme cases. Figure 4(a) shows that, when the accuracy of the altimeter instrument is very poor, large space and time windows up to 150km and 1.5h, or more, are needed in order to include a greater number of samples so that the random noises of the measurements can be reduced (see Figure 3). Conversely, if the altimeter itself is error free, the ideal treatment is of course to compare directlythe satellite and buoy measurements at the same location. The inclusion of extra samples with any spatial or temporal lag will inevitably introduce additional RMS variability (Figure 4(d)). For those altimeters with intermediate noise levels (which are likely to be the reality for the present time), the situation is more complicated. The optimal size of the window could vary from 0 to 150km in space and from 0 to 1.5h in time, depending on the competition of the wind speed variability and the altimeter noise level (Figures 4(b) and 4(c)).



Figure 4. The simulated RMS deviations of altimeter-measured wind speed as functions of spatial and temporal lags for a) J=1.2 m/s, b) J=0.8 m/s, c) J=0.4 m/s, and d) J=0.0 m/s. A buoy instrument noise of 0.5 m/s is employed in the simulation and the number of samples is 10,000 for each curve.

3.3 Influence of the wind speed variability

Now let us focus our attention on Figures 4(b) and 4(c), which can be reasonably related to the cases of Geosat and TOPEX/Poseidon, respectively. Fortunately, we are able to find some previous studies by Glazman and Pilorz (1990), Ebuchi et al. (1992) and Gower (1996) who performed qualitative or quantitative analyses of the impacts of window sizes on the comparison results using real Geosat and TOPEX/Poseidon altimeter data in conjunction with the US, Japanese and Canadian buoy data. These give us the opportunity to check the validity and applicability of our

simulation results.

The work by Ebuchi et al. (1992) is based on the analysis of 437 data pairs from the Geosat altimeter and the JMA (Japan Meteorological Agency) buoys. It is found in their results that a minimum RMS value occurs when the wind speeds are averaged between 50-100km in space or 1-1.5h in time (see Table 1 of Ebuchi et al. (1992)). In the space domain, the RMS values of the two neighboring windows, i.e., 0-50km and 100-150km, are both higher than in the central window (2.15m/s and 2.59m/s compared to 1.98m/s). This result violates the conclusion that reducing the spatial lag will persistently improve the agreement of wind speed measurements from altimeters and buoys (Hwang et al., 1998). But it supports our simulation result by coinciding with the situation between I=2 and I=3 in Figure 4(b) which represents a moderate oceanic condition. The behavior of time window in the results of Ebuchi et al. (1992) is also somewhat surprising, since the two windows with smaller temporal lags (0-0.5h and 0.5-1.0h) both exhibit higher wind speed RMS values compared to the window size usually leads to a better agreement. But again, it appears to fall into our simulation prediction between I=1 and I=2 in Figure 4(b), which may correspond to a relatively calm sea.

In another attempt by Glazman and Pilorz (1990), two space windows (0.25° and 1.0°) and two time windows (0.5h and 1.0h) are used. They conclude that the reduced spatial and temporal separations between a buoy and a satellite footprint only result in a small reduction in the observed scatter. We believe that this statement is, in a sense, representative as far as the Geosat altimeter is concerned. Since the majority of the measurements might be taken under moderate oceanic conditions as represented by *I*=2 in Figure 4(b), which almost shows a horizontal trend beyond 50km and 0.5h.

Finally, we turn to a more recent work by Gower (1996) who compares 452 pairs of TOPEX and buoy wind speed data. He uses three sizes of space window in his analysis, i.e., 107km, 20km and 10km, and finds that the RMS scatter reduces constantly with the window size from 2.05m/s to 1.76m/s. A similar trend can be found in our simulation result. In Figure 4(c), the curve of I=3 might be a good representation of Gower s case. These results imply that with the new generation of more accurate altimeters, the RMS deviation is likely to be dominated by the local wind effect. Consequently, the optimal window size might be significantly reduced in both space and time.

4 Summary and Implication

As a summary of this work, the recommended sizes of space and time windows under various options of I and J are presented in Table 2.

1								
J, I (m/s)	0	1	2	3	4			
0		0km/0h	0km/0h	0km/0h	0km/0h			
0.4	>150km/1.5h	80km/0.8h	20km/0.2h	10km/0.1h	10km/0.1h			
0.8	>150km/1.5h	140km/1.4h	100km/1.0h	50km/0.5h	20km/0.2h			
1.2	>150km/1.5h	150km/1.5h	150km/1.5h	130km/1.3h	80km/0.8h			

Table 2. The recommended sizes of space/time windows for various combinations of I and J.

With Table 2 in mind, the following points should be stressed. (1) For altimeter/buoy wind speed calibration/validation, there usually exists an optimal window size within which the RMS scatter is minimal. Under normal circumstances, this size may vary from 0 to 150km in space and from 0 to 1.5h in time. The exact value for each case depends on the noise levels of the buoy and altimeter instruments, as well as the actual spatial and temporal variabilities of the wind speed. It could be misleading to draw any general conclusion on the relationship between the size of the optimal window and the level of comparison agreement. (2) For the present generation of altimeters such as TOPEX/Poseidon, it would be ideal if a space window of 10-20km and a time window of 0.1-0.2h could be used. However, such small windows usually lead to an insufficient number of data pairs, which, in return, give rise to an increase of random errors. The eventual choice is always a comprise between the size of the window and the number of data. But as time goes on, the volume of colocation data will expand constantly, and therefore it may become a less severe constraint on the choice of windows in the future. (3) The questions raised in this study may be generalized to altimeter-derived sea level and significant wave height, or even to other geophysical parameters derived from space-borne radiometers, scatterometers and synthetic aperture radars. In this context, the results obtained here may serve as a guideline when satellite and in situ measurements are compared over the ocean.

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References

Brown, G. S., Stanley, H. R., and Roy, N. A., 1981, The wind speed measurement capability of spaceborne radar altimeters, *IEEE Journal of Oceanic Engineering*, **OE-6**, 59-63.

Carter, D. J. T., Challenor, P. G., and Srokosz, M. A., 1992, An assessment of Geosat wave height and wind speed measurements, *Journal of Geophysical Research*, **97**, 11383-11392.

Dobson, E., Monaldo, F., Goldhirsh, J., and Wilkerson, J., 1987, Validation of Geosat altimeter-derived wind speeds and significant wave heights using buoy data, *Journal of Geophysical Research*, **92**, 10719-10731.

Ebuchi, N., Kawamura, H., and Toba, Y., 1992, Growth of wind waves with fetch observed by the Geosat altimeter in the Japan Sea under winter monsoon, *Journal of Geophysical Research*, **97**, 809-819.

Glazman, R. E., and Greysukh, A., 1993, Satellite altimeter measurements of surface wind, *Journal of Geophysical Research*, **98**, 2475-2483.

Glazman, R. E., and Pilorz, S. H., 1990, Effects of sea maturity on satellite altimeter measurements, *Journal of Geophysical Research*, 95, 2857-2870.

Gower, J. F. R., 1996, Intercalibration of wave and wind data from TOPEX/POSEIDON and moored buoys off the west coast of Canada, *Journal of Geophysical Research*, **101**, 3817-3829.

Hwang, P. A., Teague, W. J., Jacobs, G. A., and Wang, D. W., 1998, A statistical comparison of wind speed, wave height, and wave period from satellite altimeters and ocean buoys in the Gulf of Mexico region, *Journal of Geophysical Research*, **103**, 10451-10468.

Monaldo, F., 1988, Expected differences between buoy and radar altimeter estimates of wind speed and significant wave height and their implications on buoy-altimeter comparisons, *Journal of Geophysical Research*, **93**, 2285-2302.

Monaldo, F., and Dobson, E., 1989, On using significant wave height and radar cross section to improve radar altimeter measurements of wind speed, *Journal of Geophysical Research*, **94**, 12699-12701.

Witter, D. L., and Chelton, D. B., 1991, A Geosat altimeter wind speed algorithm and a method for altimeter wind speed algorithm development, *Journal of Geophysical Research*, **96**, 8853-8860.