

# MAPPING LAND COVER CHANGES WITH LANDSAT IMAGERY AND SPATIO-TEMPORAL GEOSTATISTICS; AN APPLICATION TO THE PEARL RIVER DELTA, CHINA.

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## ABSTRACT

The use of remote sensing to monitor rapidly growing cities is challenging due to the dynamics of the landscape. Mapping of these land cover changes calls for an integrated algorithm that can detect both when and where changes occur but also which land covers are involved in the land conversion. The proposed method aims at improving the global observation of cities by taking into account simultaneously the spatial and the temporal relations between the various land cover types. At each pixel of the image, the time series of land cover type is modeled with transition probabilities. That time series at any specific location is estimated jointly from the local satellite information, any neighboring ground truth data, and any neighboring previously estimated time series deemed well-informed by the satellite measurements. The spatial component of the land cover types is integrated with variograms and indicator kriging.

The method is demonstrated on six Landsat images spanning nine years to monitor the growth of the city of Shenzhen, China between 1988 and 1996. The temporal prediction accuracy, i.e. all the land cover are correctly estimated at all time for a particular location, improves significantly from 33% to 61% when both spatial and temporal information are considered jointly in the estimation process.

## 1 INTRODUCTION

Mapping of land cover changes is intrinsically a space time operation, yet current image processing methods often do not consider simultaneously the spatial and temporal context when estimating land covers change.

There are two main issues with change detection: (a) the combination of the images and (b) the classification of those images. The simplest change detection method classifies each image independently, changes are then mapped by identifying which pixels have changed. The problem lies in the errors associated with the mapping of changed labels. The final accuracy is approximately the product of accuracy associated with the classification performed at each time, often with poor results.

A second method consists of analyzing the images concurrently and classifying the label trajectories. For example, instead of classifying an image into label 1 or 2, all the possible transitions between those labels (1 to 1, 1 to 2, 2 to 1 and 2 to 2) are considered, expanding to potentially  $K^{N_t}$  transitions, where  $K$  is the number of labels, and  $N_t$  the number of images.

A third method models every pixel as a time series, where the time of change is estimated. For example [Kaufmann and Seto, 2001] utilize time series econometrics to detect dates of change with better results than when the changes are obtained from post-processing independently classified images.

Another common method to process multitemporal images is the cascade approach [Swain, 1978] which consists of analyzing the sequence of image in chronological order. Past classifications being used to condition future classification.

The methodology proposed here integrates the spatial correlation of the land covers labels with temporal information, thus improving the mapping of land cover changes.

The aim is to improve existing methods of land cover change detection by considering prior knowledge about the label spatial and temporal patterns and efficiently integrating them into the classification procedure. Finally, the computational complexity of the algorithm should not increase too drastically when one increases the number of labels and/or the length of the time series.

The spatial context is integrated with geostatistics, specifically indicator kriging, which already has been used to incorporate spatial autocorrelation in estimating or simulating labels [Stein et al., 1998, Atkinson and Lewis, 2000, Brown et al., 2002, Goovaerts, 2002, Wang et al., 2004].

The framework is applied to mapping land cover changes for the city of Shenzhen, China, undergoing rapid urban growth.

## 2 NOTATIONS

Consider a domain  $D \subset \mathcal{R}^2$  sampled at different times  $t_i, i = 1, \dots, N_t \subset T$ . Let  $(\mathbf{u}, t)$  be a point in  $D \times T$  informed by a vector of length  $n_B$  of continuous attributes,  $\mathbf{Z}(\mathbf{u}, t) = \{Z_1(\mathbf{u}, t), \dots, Z_{n_B}(\mathbf{u}, t)\}$ . These attributes are the satellite measurements known as digital numbers (DN).

Each pixel  $(\mathbf{u}, t)$  must be classified into one of  $K$  labels  $\mathcal{L}_1, \dots, \mathcal{L}_K$ , for example  $K$  land cover types. Define  $I_k(\mathbf{u}, t)$  an indicator variable indicating whether or not the pixel at location  $(\mathbf{u}, t)$  has label  $\mathcal{L}_k$

$$I_k(\mathbf{u}, t) = \begin{cases} 1 & \text{if } (\mathbf{u}, t) \in \mathcal{L}_k \\ 0 & \text{otherwise} \end{cases}$$

And let

$$\mathcal{L}(\mathbf{u}, t) = k \quad \text{if } I_k(\mathbf{u}, t) = 1$$

Furthermore, let  $\Omega$  be the set of location  $\mathbf{u}_\alpha, \alpha = 1, \dots, n$  whose labels are known at all times (ground truth).  $V(\mathbf{u}, t)$

is the set of known labeled pixel data in an isochronous neighborhood of  $\mathbf{u}$  at time  $t$ .

### 3 CODING AND COMBINING INFORMATION

Following [Serpico and Melgani, 2000] and [Brown et al., 2002], three sources of information are considered relevant for the mapping land cover changes. The first and foremost source are the individual pixel satellite measurements. The second is the spatial pattern that relates land cover classes to each other. The third source of information is the temporal pattern of labels. Based on three data sources, the available information at each unsampled location  $\mathbf{u}$  is separated between isochronous (cross-sectional) and time series information. The isochronous information includes the satellite response and the neighboring land cover indicators at any specific time. The time series information consists of transition probabilities linking the land cover indicators through time. The classification at location  $\mathbf{u}$  is then done by combining these two types of information in such a way to minimize misclassification over a given training set.

#### 3.1 Information content for context

Depending of their information content, locations are separated into two groups; the well-informed time series and the poorly informed ones. Those well-informed pixels are locations where the DN measurements  $\mathbf{Z}$  alone are deemed sufficient to label them. For example, a pixel where the satellite information would indicate a probability of 0.98 or more to belong to a certain label would qualify as a well-informed node.

The information content of a time series at location ( $\mathbf{u}$ ) is measured as the sum of the maximum satellite-derived probability at each times.

$$\text{Inf}(\mathbf{u}) = \frac{1}{N_t} \sum_{i=1}^{N_T} \max(p_k^{\text{DN}}(\mathbf{u}, t_i), k = 1, \dots, K) \quad (1)$$

Those pixels, assumed to be informed adequately by the satellite information such that no spatial information is needed, are used as anchor for the less informed ones. This spreads information from high-confidence pixels to their neighbors

#### 3.2 Time series transition probabilities

Denote by  $p_k^{\text{T}}(\mathbf{u}, t)$  the probability of having label  $\mathcal{L}_k$  at location ( $\mathbf{u}, t$ ) given the co-located land cover indicators in the past ( $\mathcal{L}(\mathbf{u}, t - \Delta_1 t)$ ) and/or future ( $\mathcal{L}(\mathbf{u}, t + \Delta_2 t)$ ):

$$p_k^{\text{T}}(\mathbf{u}, t) = \Pr\{I_k(\mathbf{u}, t) = 1 \mid \mathcal{L}(\mathbf{u}, t - \Delta_1 t), \mathcal{L}(\mathbf{u}, t + \Delta_2 t)\} \quad (2)$$

The probability  $p_k^{\text{T}}(\mathbf{u}, t)$  is calibrated directly from ground truth data or determined as a function of the transition probabilities  $p_{kk'}(t_i, t_j)$  relating the probability of having class  $\mathcal{L}_{k'}$  at time  $t_j$  given that  $\mathcal{L}_k$  is observed at time  $t_i$ .

$$p_{kk'}(t_i, t_j) = \Pr\{I_{k'}(\mathbf{u}, t_j) = 1 \mid I_k(\mathbf{u}, t_i) = 1\}, \forall \mathbf{u}, k, k' \quad (3)$$

The transition probabilities  $p_{kk'}(t_i, t_j)$  are calibrated from ground truth or historical data.

#### 3.3 Isochronous probabilities

The isochronous information at any specific time is obtained by combining the satellite response and the spatial information available at that time. Denote by  $p_k^{\text{iso}}(\mathbf{u}, t)$  the isochronous probability obtained by combining the probabilities  $p_k^{\text{DN}}(\mathbf{u}, t)$  and  $p_k^{\text{S}}(\mathbf{u}, t)$  obtained from satellite and spatial information respectively.

$$\begin{aligned} p_k^{\text{iso}}(\mathbf{u}, t) &= \\ &\Pr\{I_k(\mathbf{u}, t) = 1 \mid \mathbf{Z}(\mathbf{u}, t), \mathcal{L}(\mathbf{u}', t), \mathbf{u}' \in V(\mathbf{u}, t)\} \\ &= \phi(p_k^{\text{DN}}(\mathbf{u}, t), p_k^{\text{S}}(\mathbf{u}, t)) \end{aligned} \quad (4)$$

with

$$p_k^{\text{DN}}(\mathbf{u}, t) = \Pr\{I_k(\mathbf{u}, t) = 1 \mid \mathbf{Z}(\mathbf{u}, t)\}, \forall k \quad (5)$$

and

$$p_k^{\text{S}}(\mathbf{u}, t) = \Pr\{I_k(\mathbf{u}, t) = 1 \mid \mathcal{L}(\mathbf{u}', t), \mathbf{u}' \in V(\mathbf{u}, t)\}, \forall k \quad (6)$$

where  $V(\mathbf{u}, t)$  is a set of location in the vicinity of ( $\mathbf{u}, t$ ). The combination algorithm  $\phi$  is presented later. The isochronous probability is calculated independently for each time.

**3.3.1 Satellite-derived probabilities** The conditional probability  $p_k^{\text{DN}}(\mathbf{u}, t)$  (5) for the pixel at location ( $\mathbf{u}, t$ ) to be assigned to label  $\mathcal{L}_k$  given the satellite response is computed with a classifier  $F(\cdot)$  calibrated from the known data  $\{\mathbf{Z}(\mathbf{u}_\alpha, t), \mathcal{L}(\mathbf{u}_\alpha, t)\}$  [Richards and Jia, 1999]. The function  $F(\cdot)$  approximates the conditional expectation of  $I_k(\mathbf{u}, t)$  given solely co-located satellite response. Furthermore, the classifier  $F(\cdot)$  for time  $t_i$  is calibrated only from the ground truth data available at time  $t_i$ .

In this study, the conversion of Landsat TM measurements into land cover types probabilities is done with the conventional Gaussian maximum likelihood (ML) classifier [Richards and Jia, 1999], a generative algorithm. A multiGaussian RF modeling the satellite bands is associated to each label. The probabilities  $\Pr\{I_k(\mathbf{u}, t) = 1 \mid \mathbf{Z}(\mathbf{u}, t)\}, k = 1, \dots, K$  are calculated from the training set using a Bayes' inversion

$$\begin{aligned} p_k^{\text{DN}}(\mathbf{u}, t) &= \Pr\{I_k(\mathbf{u}, t) = 1 \mid \mathbf{Z}(\mathbf{u}, t) = \mathbf{z}\} = \\ &= \frac{\Pr\{\mathbf{Z}(\mathbf{u}, t) = \mathbf{z} \mid I_k(\mathbf{u}, t) = 1\} \Pr\{I_k(\mathbf{u}, t) = 1\}}{\sum_{k'=1}^K \Pr\{\mathbf{Z}(\mathbf{u}, t) = \mathbf{z} \mid I_{k'}(\mathbf{u}, t) = 1\} \cdot \Pr\{I_{k'}(\mathbf{u}, t) = 1\}} \end{aligned}$$

Assuming the random vector  $\mathbf{Z}(\mathbf{u}, t)$  to be multiGaussian, its conditional probability is written as

$$\begin{aligned} \Pr\{\mathbf{Z}(\mathbf{u}) = \mathbf{z} \mid I_k(\mathbf{u}, t) = 1\} &= \\ &= \frac{1}{(2\pi)^{N/2} |\Sigma_i|^{1/2}} e^{-\frac{1}{2}(\mathbf{z} - \mathbf{m}_i)^T \Sigma_i^{-1} (\mathbf{z} - \mathbf{m}_i)} \end{aligned} \quad (7)$$

where  $\mathbf{m}_i$  and  $\Sigma_i$  are the mean vector and covariance matrix of the DN values belonging to the training data with label  $\mathcal{L}_i$ .

**3.3.2 Spatially-derived probabilities** Denote by  $p_k^S(\mathbf{u}, t)$  the conditional probability of observing  $\mathcal{L}_k$  at location  $(\mathbf{u}, t)$  given the isochronous label data found in the neighborhood  $V(\mathbf{u}, t)$ .

This spatial probability  $p_k^S(\mathbf{u}, t)$  may be estimated from simple indicator kriging [Goovaerts, 1997]. Simple indicator kriging is a linear interpolator that applied kriging weights to indicator data yielding the probability of belonging to a class given the neighborhood data, the marginal and the covariance model of that class.

The spatial continuity for each land cover type is measured with indicator variograms [Goovaerts, 1997] defined as

$$\gamma_k(\mathbf{h}) = E\{[I_k(\mathbf{u}, t) - I_k(\mathbf{u} + \mathbf{h}, t)]^2\}$$

The indicator kriging system uses that measure of spatial continuity to optimally assigns weight to the indicator data. The simple indicator kriging system is

$$\mathbf{K}\lambda = \mathbf{k}$$

Where  $\mathbf{K}$  is the data-to-data covariance matrix,  $\lambda$  an unknown weight vector and  $\mathbf{k}$  the unknown-to-data covariance vector. Finally the conditional probability of having label  $\mathcal{L}_k$  at location  $(\mathbf{u}, t)$  given the the neighboring data is

$$\begin{aligned} \Pr\{(\mathbf{u}, t) \in \mathcal{L}_k | I_k(\mathbf{u}', t), \mathbf{u}' \in V(\mathbf{u}, t)\} \\ = \lambda^T \mathbf{d} + (1 - \lambda^T \mathbf{1})E\{I_k(\mathbf{u}, t)\} \end{aligned} \quad (8)$$

where  $\mathbf{d}$  is the known indicator data vector and  $\mathbf{1}$  is a column vector of one. The full posterior probability density function is obtained by computing and normalizing (8) for  $k = 1, \dots, K$  such that they sum to one.

In addition to the ground truth data, the neighboring data in  $V(\mathbf{u}, t)$  also include the locations that are considered well informed by the sole satellite measurements.

The spatial context is thus accounted for through the probabilities  $p^S(\mathbf{u}, t)$  estimated with simple indicator kriging using for conditioning data the time series at those locations deemed well informed by the satellite measurements.

Taking the neighboring pixels as indicators of the presence or absence of a label assumes that those labels are somewhat related to the unknown pixel. There is, however, a risk to overextend the spatial relevance of the well-informed locations. The problem lies in the discontinuity of the landscape. For example a certain region may be predominantly urban, without forest or agriculture, yet the vegetated area could start abruptly a few pixels away. A well-informed water label located in a lake close to the shore does not say whether that shore is urbanized or vegetated, instead it tends to artificially increase the probability that the shore would belong to a water label.

To offset this problem of borders and discontinuities, the images are first segmented to find edges delineating those discontinuities. Then a data neighborhood that does not cross the edges is retained for the indicator kriging process. Interpolation (kriging) is thus limited to homogeneous neighborhoods, a schematic representation of that adaptive neighborhood is shown in Figure 1(a).

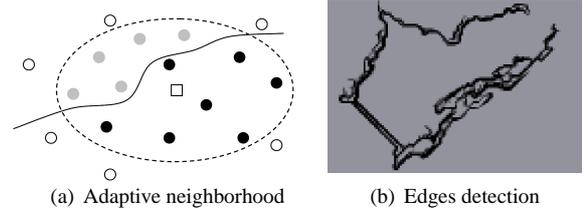


Figure 1: In (a), example of an adaptive neighborhood (dashed ellipse) in presence of an edge (solid line). The square represents the unknown location to be mapped or estimated, the empty circle are data outside the search ellipsoid, hence not considered. The grey points are data inside the search ellipsoid but on the wrong side of the edge and are not taken into account. Only the black points are considered for kriging the square location. In (b) Examples of edge detection. The edges capture the border of the bay and the dam at its extremity.

### 3.4 Posterior probability

The posterior probability  $p_k(\mathbf{u}, t)$  for class  $\mathcal{L}_k$  to occur at location  $(\mathbf{u}, t)$  is computed by combining the isochronous probability  $p_k^{iso}(\mathbf{u}, t)$  and the time series probability  $p_k^T(\mathbf{u}, t)$

$$\begin{aligned} p_k(\mathbf{u}, t) &= \Pr\{I_k(\mathbf{u}, t) = 1 \mid \text{all data}\} \\ &= \psi(p_k^{iso}(\mathbf{u}, t), p_k^T(\mathbf{u}, t)), \forall k \end{aligned}$$

The proposed combination algorithm  $\psi$  is developed in the next section.

Finally, the label  $\mathcal{L}(\mathbf{u}, t)$  is estimated by taking the most probable class of the posterior distribution:

$$\mathcal{L}^*(\mathbf{u}, t) = \arg \max_k \{p_k(\mathbf{u}, t), k = 1, \dots, K\} \quad (9)$$

The time series,  $\{\mathcal{L}(\mathbf{u}, t_1), \dots, \mathcal{L}(\mathbf{u}, t_{N_t})\}$  at location  $\mathbf{u}$  is generated with a modified cascade approach. The estimation sequence is not chronological, the time series is produced by estimating the labels starting from the best informed time, as measure from the satellite information, and then sequentially estimating the time before and after that starting time. The idea is that the starting time is very consequential for the estimation of the whole time series, that starting time is thus chosen to reduce the prediction error. If the first pixel in the time series is misclassified, it is quite likely that this misclassification will be propagated to the rest of the time series. The less informed times at any given location benefit from being conditioned on the better informed collocated times.

### 3.5 Combining probabilities

Consider the isochronous probability vector  $p^{iso}(\mathbf{u}, t)$  defined in 4 and the time series conditional probability  $p_k^T(\mathbf{u}, t)$  defined in 2 as two sources of information. Each of the those two probabilities can be transformed into a distance related to the likelihood of event  $\mathcal{L}(\mathbf{u}, t) = k$

occurring [Journal, 2002]. Let those distances be

$$x_{\mathcal{L}_k}^{\text{iso}}(\mathbf{u}, t) = \frac{1 - p_k^{\text{iso}}(\mathbf{u}, t)}{p_k^{\text{iso}}(\mathbf{u}, t)} \in [0, \infty]$$

$$x_{\mathcal{L}_k}^{\text{T}}(\mathbf{u}, t) = \frac{1 - p_{\mathcal{L}_k}^{\text{T}}(\mathbf{u}, t)}{p_{\mathcal{L}_k}^{\text{T}}(\mathbf{u}, t)} \in [0, \infty]$$

Consider also the distance related to the marginal probabilities

$$x_{\mathcal{L}_k}^{(0)} = \frac{1 - \Pr\{\mathbf{u}, t \in \mathcal{L}_k\}}{\Pr\{\mathbf{u}, t \in \mathcal{L}_k\}}, \quad \forall \mathbf{u}$$

The updated distance to the event  $\mathcal{L}(\mathbf{u}, t) = k$  occurring accounting for both information (1) and (3) is given by the ‘‘tau model’’:

$$x_{\mathcal{L}_k}(\mathbf{u}, t) = x_{\mathcal{L}_k}^{(0)} \cdot \left( \frac{x_{\mathcal{L}_k}^{\text{iso}}(\mathbf{u}, t)}{x_{\mathcal{L}_k}^{(0)}} \right)^{\tau_{\text{iso}}} \cdot \left( \frac{x_{\mathcal{L}_k}^{\text{T}}(\mathbf{u}, t)}{x_{\mathcal{L}_k}^{(0)}} \right)^{\tau_{\text{T}}} \quad (10)$$

where  $\tau_{\text{iso}}$  and  $\tau_{\text{T}}$  are parameters measuring redundancy between the two information sources [Journal, 2002, Krishnan et al., 2004]. The posterior probability is then retrieved by inverting from the updates distance (10):

$$\Pr\{\mathbf{u}, t \in \mathcal{L}_k | \mathbf{Z}(\mathbf{u}, t)\} = \frac{\frac{1}{1 + x_{\mathcal{L}_k}(\mathbf{u}, t)}}{\frac{1}{1 + x_{\mathcal{L}_1}(\mathbf{u}, t)} + \frac{1}{1 + x_{\mathcal{L}_2}(\mathbf{u}, t)} + \dots + \frac{1}{1 + x_{\mathcal{L}_K}(\mathbf{u}, t)}} \in [0, 1] \quad (11)$$

The integration of  $p^{\text{DN}}(\mathbf{u}, t)$  and  $p^{\text{S}}(\mathbf{u}, t)$  into  $p^{\text{iso}}(\mathbf{u}, t)$  is also done with expression 11 but using different tau parameters  $\tau_{\text{S}}$  and  $\tau_{\text{DN}}$ . This preliminary study assume all  $\tau$ s equal to 1, corresponding to standardized conditional independence between the sources [Krishnan et al., 2004].

The tau-model integration has convenient properties. If one of the prior probabilities is zero or one (no uncertainty), the combined probability is also zero or one. The combined probability is also always admissible, i.e. included between zero and one. The previous kriging in combination with the tau-model (11) assures that no ground truth (hard data) locations are misclassified. At any informed location  $\mathcal{L}(\mathbf{u}, t) = k$ , the kriging estimate returns a probability of one to belong to  $\mathcal{L}_k$  and zero for all others labels. The tau model then ensures that the final probability remains so.

Although mathematically similar to the logarithmic opinion pool from consensual theory [Benediktsson and Swain, 1992] and to the aggregating formula of Bordley [Bordley, 1982], the tau-model is conceptually different. The  $\tau$  exponents are not a measure of the reliability of the information source, but serves to model the redundancy between these sources. The reliability of the information is assumed to be already coded into each of the prior probability distribution related to each source. More details about the tau-model can be found in [Journal, 2002, Krishnan et al., 2004, Benediktsson and Swain, 1992].

## 4 ALGORITHMS

The algorithm proposed proceeds as follows

- Perform a segmentation of the scenes
- Compute  $p^{\text{DN}}(\mathbf{u}, t), \forall \mathbf{u}, t$ , see (5)
- Calculate the information content  $\text{Inf}(\mathbf{u}) \forall \mathbf{u}$  for all time series
- For each well-informed time series  $\mathbf{u}_\beta, \beta = 1, \dots, N_{\text{Inf}}$ 
  - Estimate time series with  $p^{\text{DN}}$ , and  $p^{\text{T}}$  (see next algorithm)
- For each remaining uninformed time series
  - Estimate time series using all information ( $p^{\text{S}}$ ,  $p^{\text{DN}}$ , and  $p^{\text{T}}$ )

The algorithm to estimate a time series at location  $\mathbf{u}$  can be described as

- Find the times  $t_i$  most informed by the satellite measurements
- Assign  $\mathcal{L}\mathbf{u}, t_i = k$  such as  $k = \arg \max_j p_j^{\text{iso}}(\mathbf{u}, t_i)$
- Estimate sequentially future times  $t_{i+1}, \dots, t_{N_i}$ 
  - Assign labels  $\mathcal{L}(\mathbf{u}, t_{i+s}), s = 1, \dots, N_t - i$  based on  $p^{\text{iso}}(\mathbf{u}, t_{i+s})$  (4) and  $p^{\text{T}}(\mathbf{u}, t_{i+s})$
- For all past times  $t_{i-1}, \dots, t_1$ 
  - Assign labels  $\mathcal{L}(\mathbf{u}, t_{i-s}), s = 1, \dots, i + 1$  based on  $p^{\text{iso}}(\mathbf{u}, t_{i-s})$  and  $p^{\text{T}}(\mathbf{u}, t_{i-s})$

## 5 A CASE STUDY OF URBANIZATION IN SHENZHEN, CHINA

This study focuses on detecting and mapping changes between 1988 and 1996 using a time series of Landsat TM images. We acquired 6 images of Shenzhen, dating from 1988, 1989, 1992, 1994, 1995 and 1996 all taken around December. The scene consists of 1 917 870 pixels approximatively covering an area of size 45km by 45 km, with each pixel of dimension 30x30 meters.

The landscape is divided into  $K=7$  classes: water, forests, agriculture, urban, fish pond, transition (land getting cleared for urban settlement) and shrub. The ground truth measurements consists of 1917 locations identified by expert interpretation or by field reconnaissance. At ground truth locations the labels are deemed known at all times. The prediction errors, the expected errors between the estimated label and the true label at any location, are estimated by a 5-fold cross-validation procedure [Hastie and Friedman, 2001]. The known labels are divided five times, each time into a training set and a testing set such that all samples are used once for testing purposes. Each split is done such that 80% of the ground truth data belong to the training set and the remainder 20% to the test set.

## 5.1 Results

The results of the proposed method applied to the Shenzhen scene are compared to the accuracy resulting from the maximum likelihood (ML) classifier, see expression 7. The ML classification is done by assigning to a time-space location  $(\mathbf{u}, t)$ , the class that has the maximum probability  $p_k^{\text{DN}}(\mathbf{u}, t)$ . This ML classification considers only the satellite responses thus ignores the temporal and spatial correlation between labels. The changes are mapped by comparing the ML classification performed independently at each time.

The results are validated using (a) an overall accuracy criterion, the percent of correctly classified pixels, and (b) a time series accuracy, the percent of locations which have their vector of labels **all** correct. A time series at location  $\mathbf{u}$  is well classified only if its six labels have been correctly predicted. For change detection purposes, the time series accuracy is critical as it measures how well the changes are mapped in both time and space. This allows for a correct identification of the time when the landscape has changed, and from which label to which other label. Such information is necessary to study the dynamics of the landscape.

With the ML classifier, the accuracy from a five-fold cross validation yields an overall accuracy of 78%, but the temporal accuracy is only 33%. The proposed method marginally improves the overall accuracy from 78% to 82%. However, the temporal accuracy goes up to 61%, a considerable improvement.

Tables 1 and 2 show the confusion matrices of the classification using ML and then using the proposed method which utilizes both spatial and temporal information. The producer and consumer accuracy for both cases are shown in Table 3. For all labels the consumer accuracy is greater or at least equal when space and time are incorporated in the classification. The improvement is especially noticeable for the forest class, the consumer accuracy jump from 44% to 91% as less forest were misclassified into shrub. The producer accuracy, however, remains low with both methods for the forest class. The shrub ground truth data are almost all correctly identified (a producer accuracy of 0.93), a significant improvement from 0.83 when only the ML classifier is used. The fish ponds class also greatly benefits from the proposed method as its consumer accuracy goes from 0.66 to 0.81 and its producer accuracy improves 20 points from 0.72 to 0.92. The proposed method reduced the misclassification of fish pond into the water class from 56 misclassifications to only 6. Only the producer accuracy of water and forest decreases with the integration of spatial and temporal information.

The indicator kriging step decreases the level of speckling in the images, producing smoother maps. For example, the ML tend to classify many shadow zones in mountainous areas as water; the integration of spatial information corrects many of those misclassified pixels. There is no need to post-process the classified images to remove the speckles.

labels	wat.	for.	agr.	urb.	pond	tra.	shr.
wat.	1584	9	18	12	60	40	9
for.	27	185	12	0	0	0	287
agr.	35	23	988	39	0	24	503
urb.	4	2	30	877	5	292	55
pond	56	0	1	2	165	3	2
tra.	36	1	23	257	9	2535	77
shr.	13	204	215	44	12	57	2670

Table 1: Confusion table for maximum likelihood

labels	wat.	for.	agr.	urb.	pond	tra.	shr.
wat.	1503	5	24	81	35	40	44
for.	23	168	26	7	0	0	287
agr.	39	6	1059	15	0	49	444
urb.	0	3	23	938	0	265	36
pond	6	1	5	1	210	4	2
tra.	38	0	23	239	10	2569	59
shr.	6	2	116	34	3	60	2991

Table 2: Confusion table with space/time consideration

Maximum likelihood							
	wat.	for.	agr.	urb.	pond	tra.	shr.
C. Acc	0.9	0.44	0.77	0.71	0.66	0.86	0.74
P. Acc	0.91	0.36	0.61	0.69	0.72	0.86	0.83
Proposed method							
C. Acc	0.93	0.91	0.83	0.71	0.81	0.86	0.77
P. Acc	0.87	0.33	0.65	0.74	0.92	0.87	0.93

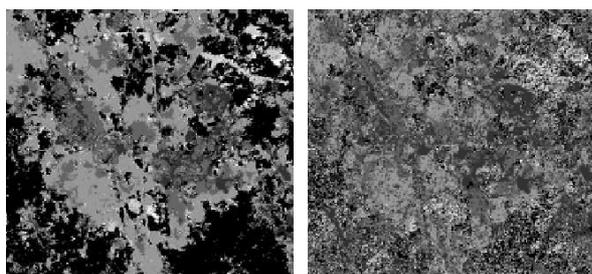
Table 3: Consumer and producer accuracy for both methods

Kriging based on well-informed locations with an adaptive neighborhood has some advantages over commonly used post-processing filters such as the majority filter. With that filter any pixel can be changed according to the neighboring pixels; hence there is a possibility that a pixel well informed by the satellite measurement, thus with a low probability of misclassification, gets changed based on potentially more uncertain pixels in its neighborhood. Such scenario cannot happen with our proposed kriging as the spatial context is created from those well-informed pixels and propagated to the poorly informed ones, never the other way.

The maps in Figure 2 show, for each location, the year at which change first occurred. A comparison between Figure 2(a) and (b) shows that the proposed method preserves some spatial relationships for the land cover changes, exhibiting a structured evolution of the landscape. On the contrary, the ML method produces a salt and pepper texture where the physical evolution of the landscape is difficult to discern.

The proposed method considerably reduced the number of false positives. With the ML prediction, 35% of locations had changed more than once, a number that expert visual inspection of the images does not validate. With the proposed method only 9% of locations are predicted to change more than once. The ML also predicts that 22% of locations did not change while that percent goes up to 64%

when both spatial and temporal information are combined for prediction.



(a) Year of first changes for proposed method (b) Year of first changes for ML method

Figure 2: Map of predicted land cover changes representing the year at which the first change occurred. Figure (a) maps the year of change as predicted by the proposed method. Figure (b) does it for the ML method. Black indicates no changes, lighter tones indicates later times. Note the greater spatial resolution for the proposed method.

## 6 CONCLUSION

This paper presents a new methodology that integrates the spatial and temporal autocorrelation of labels in remote sensing applications. That integration results in a more accurate change detection map that better identifies what, when, and where landscape changes occurred. This study uses only 6 images, the extension to longer time series would be straightforward as the complexity of the algorithm only increases linearly with additional images.

Prior identification of well-informed locations from only satellite information is shown to work well.

The algorithm is flexible as it can handle any spatial, temporal or satellite classifier as long as it provides the probability for a pixel to belong to any label. For example, the ML classifier could be replaced without any change to the algorithm by a probabilistic neural networks or any other suitable algorithm. In the same way, a Gibbs-Markov RF or multiple points geostatistical algorithm [Strebelle, 2002] could replace the indicator kriging to provide the spatial context. Furthermore, if integrated with a GIS, other sources of information ( such as distance to road, topography) could be integrated if that information can be coded into probabilities. Optimizing the tau exponents through some kind of training algorithm could further increase the classification accuracy.

Most importantly, the new method shows a considerable increase in the temporal accuracy with the proposed method. The evolution of the landscape displays greater spatial continuity and appears more realistic.

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